Mortality of hospitalised internal medicine patients bedspaced to non-internal medicine inpatient units: retrospective cohort study

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ABSTRACT

Objective To compare inhospital mortality of general internal medicine (GIM) patients bedspaced to off-service wards with GIM inpatients admitted to assigned GIM wards.

Method A retrospective cohort study of consecutive GIM admissions between 1 January 2015 and 1 January 2016 was conducted at a large tertiary care hospital in Canada. Inhospital mortality was compared between patients admitted to off-service wards (bedspaced) and assigned GIM wards using a Cox proportional hazards model and a competing risk model. Sensitivity analyses included propensity score and pair matching based on GIM service team, workload, demographics, time of admission, reasons for admission and comorbidities.

Results Among 3243 consecutive GIM admissions, more than a third (1125, 35%) were bedspaced to off-service wards with the rest (2118, 65%) admitted to assigned GIM wards. In hospital, 176 (5%) patients died: 88/1125 (8%) bedspaced patients and 88/2118 (4%) assigned GIM ward patients. Compared with assigned GIM wards patients, bedspaced patients had an HR of 3.42 (95% CI 2.23 to 5.26; P<0.0001) for inhospital mortality at admission, which then decreased by HR of 0.97 (95% CI 0.94 to 0.99; P=0.0133) per day in hospital. Competing risk models and sensitivity analyses using propensity scores and pair matching yielded similar results.

Conclusions Bedspaced patients had significantly higher inhospital mortality than patients admitted to assigned GIM wards. The risk was highest at admission and subsequently declined. The results of this single centre study may not be generalisable to other hospitals and may be influenced by residual confounding. Despite these limitations, the relationship between bedspacing and patient outcomes requires investigation at other institutions to determine if this common practice represents a modifiable patient safety indicator.

INTRODUCTION

Hospital capacity strain negatively impacts quality of care.1 Overcrowding creates many challenges to efficient patient navigation of the hospital system from initial assessment in the emergency department (ED), to admission onto inpatient wards, and ultimately to patient discharge. Beginning in the ED, overcrowding is associated with delayed care and reduced survival.2 After admission, patients may then become ‘boarders’, who remain in EDs after admission while awaiting inpatient beds. ED boarding delays medical care, prolongs length of stay and may increase mortality.3–5

When assigned general internal medicine (GIM) ward beds become fully occupied, new patients are often ‘bedspaced’, meaning they are admitted to off-service inpatient units to free up ED beds.7–8 This is a common occurrence across hospitals.9 Bedspaced patients reside in areas of the hospital away from assigned GIM wards where medical team members responsible for their care are usually located. Allied health services (eg, nursing, physiotherapy) for bedspaced patients are typically delivered by the ‘host service’, and they may have limited contact and communication with the GIM team responsible for patient care.8,9

Little research exists to assess whether bedspacing compromises quality of care or patient outcomes. In other areas of medicine, patients who are not cared for in the appropriate specialised setting may have worse quality of care and outcomes; for example, stroke patients not on specialised stroke services9 and haemodialysis patients on surgical services.10 Similar issues may apply to bedspaced GIM patients.

This study aimed to evaluate the relationship between bedspacing of GIM patients and inhospital mortality.

METHODS

Design

We conducted a retrospective cohort study using data collected during 2015...
at a large Canadian tertiary care hospital. Ethical approval was obtained from the institution’s research ethics committee. We based our study reporting on the REporting of studies Conducted using Observational Routinely-collected health Data (RECORD) guidelines for observational studies using routinely collected data.11

Setting
A tertiary level, acute-care, university-affiliated teaching hospital with a large catchment area exceeding 500,000 people in Canada served as the study setting. The hospital has 440 inpatient beds and admits more than 22,000 inpatients each year.

The GIM service has 68 beds for four clinical teaching unit (CTU) teams, which consist of a staff physician, a third year internal medicine resident, a second year internal medicine resident, two to three first year residents and two to three medical student clerks. The CTU teams change every 4 weeks. Staff physicians on CTU teams rotate every 2 weeks.

Patient population
We included consecutive adult patients admitted to the four GIM CTU teams at the hospital from 1 January 2015 to 1 January 2016. Patients were excluded if they had any of the following:
▶ Admission to the medical short stay team (anticipated length of stay <72 hours)
▶ Located in the step down intensive care unit
▶ Transfer to another service during hospital admission
▶ Death, discharge or leaving against medical advice before being transferred to a ward bed

Data sources
Data were obtained from patient electronic medical records and the Discharge Abstract Database, an administrative database collected for all hospital discharges and maintained by the Canadian Institute of Health Information (CIHI).12 Data were collected on patient demographics, admission diagnosis, comorbidities, death in hospital and details of discharge from hospital.

Admission and bedspacing
Patients were admitted under one of four CTU teams. Patient assignment to CTU teams was based on day of admission.

On admission, patients were typically transferred from the ED to the GIM ward. When admission numbers exceeded assigned GIM ward bed capacity, the patient flow manager secured a bed on an off-service ward based on services with available beds (with medical subspecialty wards given priority over surgical wards) and length of time the patient was in the ED since admission. In principle according to hospital guidelines, patients’ reasons for admission, severity of illness, comorbidities, CTU team assignment, plans of care in hospital and discharge dispositions were not considered in this decision.

The two assigned GIM wards with 68 beds were located on geographically consolidated hospital wings (wards A and B). If all beds were occupied on the assigned GIM wards, patients would be bedspaced to off-service ward beds. These patients were defined as bedspaced patients.

The medicine off-service wards included:
C: Medicine Short Stay Unit, Alternate Level of Care patients
D: Haematology/Oncology, Medical Oncology, Radiation Oncology, Palliative Care
H: Cardiac Surgery, Thoracic Surgery, Vascular Surgery
E: Gynaecology Oncology, Plastic Surgery, Urology
G: Otolaryngology, General Surgery
J: Orthopaedic Surgery

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ICD-10-CA diagnosis codes present at and after admission for each patient were used to calculate the Charlson Comorbidity Index (CCI) using an established coding algorithm.

Outcomes
The primary outcome was all-cause inhospital mortality. Patients were followed until death in hospital, discharge, or for up to 365 days following hospital admission, whichever came first. Length of stay was calculated as time from admission to discharge or death in hospital. Readmissions to the same hospital within 30 days were recorded for patients who were discharged alive.

Statistical analysis
A Cox proportional hazards model was used to characterise inhospital mortality. The survival analysis assumed non-informative censoring. The prognosis of patients who were discharged may be different than that of patients who remained hospitalised. The potential impact of informative censoring was addressed in a competing risk model that accounted for competing risk events that would preclude the event of interest.

Specifically, as a secondary analysis, a competing risk model was used to describe inhospital mortality and time to discharge, where possible end points included being alive in hospital on day 365, dying in hospital or being discharged alive. Based on a cumulative incidence function, a subdistribution HR (sHR) was calculated based on the Fine and Gray model. Assumption of proportional hazard was tested based on weighted Schoenfeld residuals for the Cox proportional hazards and competing risk models.

To account for non-proportional hazards, an interaction of bedsparing variable with time in a linear relationship was added in the Cox proportional hazards and competing risk models.

As sensitivity analyses, a propensity score for room allocation (bedspaced vs assigned GIM ward) was estimated using a multivariable logistic regression model. The following covariates were included in the model: age, sex, admission month, admission on a weekend, capacity ratio, cumulative CCI score, individual CCI comorbidity, RIW and CMG. CMG had a factor for each of the 52 most common groups, which accounted for approximately 80% of all admissions. Bedspaced patients were matched in a 1:1 ratio to assigned GIM ward patients using nearest neighbour matching with a specified calliper width of 0.2 times SD of the logit of the propensity score. Covariate balance between the two groups was assessed by standardised difference with a threshold of 0.10 for significant imbalance. The two groups matched by propensity scores were compared using the Cox proportional hazards and competing risk models.

As another sensitivity analysis, bedspaced patients were matched to assigned GIM ward patients in a 1:1 ratio, so that each matched pair was admitted in the same 4-week rotation block under the same CTU team with the same CMG. This ensured the same CTU team for each matched pair. In cases where multiple matched pair combinations were possible, the matched pair with the closest admission dates was selected. Thus, each matched pair was admitted under the same CTU team at a similar time to reflect similar workload. The two groups were compared with the Cox proportional hazards and competing risk models.

Our study data were complete without any missing data, so statistical handling of missing data was not applicable.

All reported CIs were two-sided 95% intervals and all tests were two-sided with a P<0.05 significance level. All analyses were done with R V.3.2.3 (R Foundation for Statistical Computing, Vienna, Austria). Matching was done using the package matchit.

Competing risk analysis was done using the packages cmprsk and crrSC.

RESULTS
Patient characteristics
From 1 January 2015 to 1 January 2016, 4925 patients were admitted under the GIM service, of which 3243 patients were included in the analysis (figure 1). Of these 3243 patients, over a third (1125, 35%) were bedspaced to off-service wards while the rest (2118, 65%) were admitted to assigned GIM wards (table 1 and online supplementary appendix table 1). Numbers of bedspaced patients admitted under each off-service ward are listed in online supplementary appendix table 2.

Bed-spacing
Over the 1-year period, the daily proportion of bedspaced patients had a median of 32% (IQR 28%–35%), minimum of 19% and maximum of 47% (online supplementary appendix figure 1). Daily numbers of GIM inpatients were consistently over the capacity of 68 patients except for 5 days in May (online supplementary appendix figure 2).

Bed moves were uncommon and occurred equally in both directions. Only 365 (11%) patients were on different wards at time of discharge or death than their admission wards: 195 (6%) were originally bedspaced and then transferred to assigned GIM wards, while 170 (5%) were originally on assigned GIM wards and then bedspaced to off-service wards.

Outcomes
Of 3243 GIM inpatients, 176 (5%) patients died in hospital including 88/1125 (8%) bedspaced patients and 88/2118 (4%) assigned GIM wards patients. Inhospital mortality rates for individual wards and each CCI strata are listed in online supplementary appendix tables 2 and 3, respectively. Deaths occurring among
assigned GIM ward and bedspaced patients are illustrated in online supplementary appendix figure 3.

Of the 3067 patients who were discharged alive, 205/1037 (20%) bedspaced patients and 372/2030 (18%) assigned GIM wards patients were readmitted to the same hospital within 30 days (P=0.3582).

In an unadjusted Cox proportional hazards model, bedspaced patients had significantly greater in-hospital mortality than assigned GIM wards patients (HR 2.24; 95% CI 1.67 to 3.02; P<0.0001). Testing for non-proportionality was significant (P=0.0040), indicating that the relationship between bedspacing and death was time-dependent. After adding an interaction term of bedspacing with time, the HR for mortality of bedspaced patients was 3.42 (95% CI 2.23 to 5.26; P<0.0001) at admission, which then decreased by an HR of 0.97 (95% CI 0.94 to 0.99; P=0.0133) per day. The HRs for time periods of 0–7 days, 7–14 days, 14–21 days, and 21–28 days were 3.26 (95% CI 2.10 to 5.07), 2.68 (95% CI 1.39 to 5.16), 1.43 (95% CI 0.53 to 3.86) and 1.16 (95% CI 0.40 to 3.34), respectively.

Cox proportional hazards models were used to estimate the effects of bedspaced patients in GIM wards on mortality and discharge status. In an unadjusted competing risk model, bedspaced patients had a significantly higher risk of inhospital mortality than assigned GIM wards patients (HR 2.42; 95% CI 1.67 to 3.50; P<0.0001). Testing for non-proportionality was significant (P=0.0040), indicating that the relationship between bedspacing and death was time-dependent. After adding an interaction term of bedspacing with time, the HR for mortality of bedspaced patients was 3.18 (95% CI 2.03 to 4.98; P<0.0001) at admission, which then decreased by an sHR of 0.96 (95% CI 0.93 to 0.99; P=0.0124) per day. The sHR of bedspaced patients for being discharged alive was 1.14 (95% CI 1.02 to 1.26; P=0.0210) at admission, which then decreased by an sHR of 0.98 (95% CI 0.98 to 0.99; P=0.0002) per day.

### Propensity score matching

Based on propensity score matching, 1093 bedspaced patients were matched to 1093 assigned GIM wards patients (table 2 and online supplementary appendix table 4). The maximum standardised difference of matched variables was 0.0579 (5.79%), reflecting good balance on measured baseline characteristics.

In the propensity-matched sample, 138/2186 (6%) patients died in hospital: 87/1093 (8%) for bedspaced patients and 51/1093 (5%) for assigned GIM wards patients.

In a Cox proportional hazards model using the propensity-matched sample, the HR for mortality of bedspaced patients was 3.68 (95% CI 2.30 to 5.66; P<0.0001) at admission, which then decreased by an HR of 0.95 (95% CI 0.93 to 0.98; P<0.0001) per day. The sHR for discharged alive of bedspaced patients was 1.07 (95% CI 1.01 to 1.13; P=0.0220) at admission, which then decreased by an sHR of 0.98 (95% CI 0.97 to 0.99; P=0.0002) per day.

In a competing risk model using the propensity-matched sample, bedspaced patients had an sHR for inhospital mortality of 2.83 (95% CI 1.67 to 4.79; P<0.0001) at admission, which then decreased by an sHR of 0.96 (95% CI 0.93 to 1.00; P=0.0302) per day. The sHR for discharged alive of bedspaced patients was 1.13 (95% CI 1.00 to 1.27; P=0.0530).
Pair matching
Matching by 4-week rotation blocks, CTU team and CMG, 456 bedspaced patients were matched to 456 assigned GIM wards patients (table 3 and online supplementary appendix table 5). Although the two groups were not matched by comorbidities, the proportions of patients with CCI comorbidities were similar between the two groups. The maximum standardised difference of comorbidities was 0.1087. Among matched pairs, the median difference in days between admission dates was 5.00 days (IQR 2.00–11.00 days). Of the 456 matched pairs, 168 (37%) pairs were admitted in the same week.

In the paired matching sample, 50/912 (5%) patients died in hospital: 36/456 (8%) bedspaced patients and 14/456 (3%) assigned GIM wards patients.

In a Cox proportional hazards model, compared with assigned GIM wards patients, the HR for mortality of bedspaced patients was 3.84 (95% CI 1.42 to 10.38; P=0.0080) at admission, which then decreased by an HR of 0.98 (95% CI 0.92 to 1.05; P=0.5485) per day.

### Table 1 Baseline characteristics of bedspaced and assigned general internal medicine (GIM) wards patients

<table>
<thead>
<tr>
<th>Demographics</th>
<th>All patients (n=3243)</th>
<th>Bedspaced patients (n=1125)</th>
<th>Assigned GIM ward patients (n=2118)</th>
<th>Bedspaced versus assigned GIM standardised difference</th>
</tr>
</thead>
<tbody>
<tr>
<td>Male</td>
<td>1631 (50.3%)</td>
<td>578 (51.4%)</td>
<td>1053 (49.7%)</td>
<td>0.0332</td>
</tr>
<tr>
<td>Age median (IQR)</td>
<td>71.00 (59.00–81.00)</td>
<td>71.00 (57.00–81.00)</td>
<td>72.00 (59.00–82.00)</td>
<td>0.1010</td>
</tr>
<tr>
<td>Length of stay</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Length of stay, days: median (IQR)</td>
<td>5.72 (3.38–10.91)</td>
<td>5.31 (3.00–9.65)</td>
<td>5.97 (3.48–11.80)</td>
<td>0.1119</td>
</tr>
<tr>
<td>Admission day</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Weekend (Saturday or Sunday)</td>
<td>838 (25.8%)</td>
<td>317 (28.2%)</td>
<td>521 (24.6%)</td>
<td>0.0813</td>
</tr>
<tr>
<td>Admission month</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>January</td>
<td>271 (8.4%)</td>
<td>107 (9.5%)</td>
<td>164 (7.7%)</td>
<td>0.0630</td>
</tr>
<tr>
<td>February</td>
<td>231 (7.1%)</td>
<td>82 (7.3%)</td>
<td>149 (7.0%)</td>
<td>0.0098</td>
</tr>
<tr>
<td>March</td>
<td>282 (8.7%)</td>
<td>85 (7.6%)</td>
<td>197 (9.3%)</td>
<td>0.0629</td>
</tr>
<tr>
<td>April</td>
<td>246 (7.6%)</td>
<td>69 (6.1%)</td>
<td>177 (8.4%)</td>
<td>0.0859</td>
</tr>
<tr>
<td>May</td>
<td>256 (7.9%)</td>
<td>83 (7.4%)</td>
<td>173 (8.2%)</td>
<td>0.0295</td>
</tr>
<tr>
<td>June</td>
<td>291 (9.0%)</td>
<td>111 (9.9%)</td>
<td>180 (8.5%)</td>
<td>0.0474</td>
</tr>
<tr>
<td>July</td>
<td>297 (9.2%)</td>
<td>96 (8.5%)</td>
<td>201 (9.5%)</td>
<td>0.0334</td>
</tr>
<tr>
<td>August</td>
<td>301 (9.3%)</td>
<td>91 (8.1%)</td>
<td>210 (9.9%)</td>
<td>0.0638</td>
</tr>
<tr>
<td>September</td>
<td>250 (7.7%)</td>
<td>104 (9.2%)</td>
<td>146 (6.9%)</td>
<td>0.0864</td>
</tr>
<tr>
<td>October</td>
<td>281 (8.7%)</td>
<td>104 (9.2%)</td>
<td>177 (8.4%)</td>
<td>0.0313</td>
</tr>
<tr>
<td>November</td>
<td>266 (8.2%)</td>
<td>94 (8.1%)</td>
<td>172 (8.4%)</td>
<td>0.0085</td>
</tr>
<tr>
<td>December</td>
<td>271 (8.4%)</td>
<td>99 (8.8%)</td>
<td>172 (8.1%)</td>
<td>0.0244</td>
</tr>
<tr>
<td>Capacity ratio median (IQR)</td>
<td>1.34 (1.25–1.41)</td>
<td>1.35 (1.26–1.41)</td>
<td>1.34 (1.25–1.41)</td>
<td>0.0396</td>
</tr>
<tr>
<td>Clinical teaching unit (CTU) team</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>A</td>
<td>820 (25.3%)</td>
<td>264 (23.5%)</td>
<td>556 (26.3%)</td>
<td>0.0644</td>
</tr>
<tr>
<td>B</td>
<td>786 (24.2%)</td>
<td>248 (22.0%)</td>
<td>538 (25.4%)</td>
<td>0.0790</td>
</tr>
<tr>
<td>C</td>
<td>827 (25.5%)</td>
<td>315 (28.0%)</td>
<td>512 (24.2%)</td>
<td>0.0872</td>
</tr>
<tr>
<td>D</td>
<td>810 (25.0)</td>
<td>298 (26.5%)</td>
<td>512 (24.2%)</td>
<td>0.0533</td>
</tr>
<tr>
<td>Charlson Comorbidity Index Score</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>0</td>
<td>888 (27.4%)</td>
<td>324 (28.8%)</td>
<td>564 (26.6%)</td>
<td>0.0485</td>
</tr>
<tr>
<td>1</td>
<td>708 (21.8%)</td>
<td>236 (21.0%)</td>
<td>472 (22.3%)</td>
<td>0.0318</td>
</tr>
<tr>
<td>&gt;=2</td>
<td>1647 (50.8%)</td>
<td>565 (50.2%)</td>
<td>1082 (51.1%)</td>
<td>0.0173</td>
</tr>
<tr>
<td>Resource intensity weight (RIW) Median (IQR)</td>
<td>1.08 (0.70–1.90)</td>
<td>1.04 (0.69–1.72)</td>
<td>1.11 (0.71–2.01)</td>
<td>0.0757</td>
</tr>
<tr>
<td>Top seven case mix group (admission diagnosis)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>139 Chronic obstructive pulmonary disease</td>
<td>313 (9.7%)</td>
<td>97 (8.6%)</td>
<td>216 (10.2%)</td>
<td>0.0540</td>
</tr>
<tr>
<td>138 Viral/unspecified pneumonia</td>
<td>188 (5.8%)</td>
<td>62 (5.5%)</td>
<td>126 (6.0%)</td>
<td>0.0188</td>
</tr>
<tr>
<td>487 Lower urinary tract infection</td>
<td>152 (4.7%)</td>
<td>37 (3.3%)</td>
<td>115 (5.4%)</td>
<td>0.1050</td>
</tr>
<tr>
<td>196 Heart failure without cardiac catheter</td>
<td>125 (3.9%)</td>
<td>42 (3.7%)</td>
<td>83 (3.9%)</td>
<td>0.0097</td>
</tr>
<tr>
<td>477 Renal failure</td>
<td>110 (3.4%)</td>
<td>41 (3.6%)</td>
<td>69 (3.3%)</td>
<td>0.0212</td>
</tr>
<tr>
<td>254 Gastrointestinal haemorrhage</td>
<td>88 (2.7%)</td>
<td>41 (3.6%)</td>
<td>47 (2.2%)</td>
<td>0.0846</td>
</tr>
<tr>
<td>437 Diabetes</td>
<td>76 (2.3%)</td>
<td>25 (2.2%)</td>
<td>51 (2.4%)</td>
<td>0.0123</td>
</tr>
</tbody>
</table>

For full list of Charlson Comorbidity Index comorbidities and top 52 case mix groups (CMGs), please see online supplementary appendix table 1.
In a competing risk model, bedspaced patients had an sHR for inhospital mortality of 3.73 (95% CI 1.38 to 10.09; P=0.0096) at admission, which then decreased by an sHR of 0.97 (95% CI 0.91 to 1.04; P=0.3589) per day. The sHR of bedspaced patients for being discharged alive was 1.11 (95% CI 0.90 to 1.37; P=0.3365) at admission, which decreased by an sHR of 0.98 (95% CI 0.96 to 1.00; P=0.0477) per day.

Sensitivity analyses

Comparison of bedspaced and assigned GIM wards patients using restricted mean time survival that did not assume proportional hazards again demonstrated shorter survival time for bedspaced patients (online supplementary appendix text 1).

Comparison of bedspaced and assigned GIM wards patients based on their first ward bed (instead of ward bed at discharge or death) yielded similar results (online supplementary appendix table 6).

An analysis of matched pairs admitted in the same week (to ensure the same staff physician for both matched patients) showed similar results (online supplementary appendix table 7).

**DISCUSSION**

In this retrospective cohort study, the risk for inhospital mortality among bedspaced patients was approximately three times that of patients admitted to assigned GIM wards in the first week of admission, which then decreased to approximately the same risk by the third week. This result was consistent across multiple analyses that employed competing risk models, propensity score matching and pair matching.

To our knowledge, this is the first study to document heightened mortality associated with the practice of bedspacing.

The CCI Score and length of stay in our study were lower than those reported in some individual tertiary Canadian centres, but higher than the average reported in a multicentre study of 11 hospitals in Ontario. Our inhospital mortality rate of 5%...
approximates the inhospital mortality rate on GIM services in Canadian academic tertiary acute care hospitals.12

This study has several strengths. First, the large sample size allowed for precise estimates. Second, our study recruited patients over a 1-year period. The admitting diagnoses, prognoses and outcomes may differ depending on month and season. Data over a 1-year period allowed examination and adjustment for trends over time. Third, the findings remained robust using Cox proportional hazards and competing risk models with adjustment for many potential confounders using propensity score matching and pair matching. The propensity score and pair matching ensured that the two groups were balanced in terms of prognostic factors including demographics, comorbidity, care provider, workload, time of admission and reason for admission.

Several study limitations merit emphasis. First, this study used data from a retrospective discharge database. The database was not created specifically for our research question. However, the data collection and verification were rigorous, consistent and complete, as it was based on a standardised protocol by CIHI.12 All of the data relevant to our research were collected by the database.

Second, there may be residual confounders. As an example, bedspacing may be more common during periods of higher patient influx, when the heavier caseload may have led to worse outcomes. However, the temporal trend showed a constant trend in patient volume above capacity of 68 beds and the proportion of bedspaced patients with few significant spikes (online supplementary appendix figures 1 and 2). Moreover, we adjusted for workload by including a capacity ratio in the propensity score and by pair matching (where each matched pair was admitted at a similar time under the same CTU team). Boarding time in the emergency room may be another confounder. Patients with longer boarding time may be more likely to be bedspaced to off-service wards, where the boarding time led to a higher mortality. We excluded patients who were discharged or died in the emergency room. Thus, all patients were alive on completion of transfer to the hospital ward. More than 85% of inhospital deaths occurred after 2 days in hospital. Since all patients were alive after transfer to the hospital ward and the majority of deaths occurred after 2 days of hospital stay, it would be unlikely for cause of death to be due to care during boarding in the emergency room and more likely for the cause of death to be related to the care received in the hospital ward. Another possible confounder would be preferential bedspacing of sicker patients, which could account for higher mortality. However, this was unlikely given that bedspacing decisions were not based on disease severity. Although disease severity was not adjusted, the comorbidity score and reason for admission were similar between the propensity-matched groups. Our primary analysis did not account for transfer of patients between off-service and assigned GIM wards. However, transfers occurred in only 11% of patients and went in both directions. Sensitivity analysis based on first ward bed yielded similar results (online supplementary appendix table 6). Transfer between the off-service and GIM wards would make the two groups more similar to one

Table 3  Bedspaced and assigned general internal medicine (GIM) wards patients pair matched by rotation block, Clinical teaching unit (CTU) team and case mix group (CMG)  

<table>
<thead>
<tr>
<th>Demographics</th>
<th>Bedspaced patients (n=456)</th>
<th>Assigned GIM wards patients (n=456)</th>
<th>Standardised difference</th>
</tr>
</thead>
<tbody>
<tr>
<td>Male</td>
<td>232 (50.9%)</td>
<td>222 (48.7%)</td>
<td>0.0439</td>
</tr>
<tr>
<td>Age median (IQR)</td>
<td>73.00 (61.00–82.25)</td>
<td>73.00 (61.00–83.00)</td>
<td>0.0393</td>
</tr>
<tr>
<td>Length of stay</td>
<td>4.97 (2.96–8.84)</td>
<td>5.54 (3.31–9.98)</td>
<td>0.0774</td>
</tr>
</tbody>
</table>

Admission day
Weekend (Saturday or Sunday)  123 (27.0%) 112 (24.6%) 0.0552

Admission month
January  35 (7.7%) 37 (8.1%) 0.0163
February  41 (9.0%) 31 (6.8%) 0.0814
March  35 (7.7%) 46 (10.1%) 0.0849
April  33 (7.2%) 32 (7.0%) 0.0085
May  29 (6.4%) 27 (5.9%) 0.0183
June  37 (8.1%) 37 (8.1%) 0.0000
July  45 (9.9%) 46 (10.1%) 0.0073
August  34 (7.3%) 39 (8.6%) 0.0404
September  43 (9.5%) 37 (8.1%) 0.0614
October  39 (8.6%) 45 (9.9%) 0.0455
November  45 (9.9%) 39 (8.6%) 0.0455
December  38 (8.3%) 40 (8.8%) 0.0157

Charlson Comorbidity Index (CCI) Score
>=2  222 (48.7%) 237 (52.0%) 0.0458

Resource intensity weight (RIW) Median (IQR)
0  1.35 (1.26–1.41) 1.34 (1.26–1.43) 0.0346

CTU team
A  122 (26.8%) 122 (26.8%) 0.0000
B  107 (23.5%) 107 (23.5%) 0.0000
C  113 (24.8%) 113 (24.8%) 0.0000
D  114 (25.0%) 114 (25.0%) 0.0000

Top seven case mix group (admission diagnosis)
139 Chronic obstructive pulmonary disease  83 (18.2%) 83 (18.2%) 0.0000
138 Viral/unspecified pneumonia  52 (11.4%) 52 (11.4%) 0.0000
487 Lower urinary tract infection  32 (7.0%) 32 (7.0%) 0.0000
196 Heart failure without cardiac catheter  31 (6.8%) 31 (6.8%) 0.0000
477 Renal failure  23 (5.0%) 23 (5.0%) 0.0000
254 Gastrointestinal haemorrhage  20 (4.4%) 20 (4.4%) 0.0000
437 Diabetes  15 (3.3%) 15 (3.3%) 0.0000

For full list of Charlson Comorbidity Index comorbidities and top 52 case mix groups (CMGs), please see online supplementary appendix table 7.
another making it more difficult to detect a difference, so results would be conservative.

Third, our study was based on a single academic acute care hospital, which limits the generalisability of the results.

There may be several reasons for bedspaced patients to have a higher risk of mortality. First, off-service wards were geographically separated from the CTU teams’ usual rounding area, which may have decreased the frequency and duration of patient contact and interaction. Decreased patient contacts in situations such as boarding in EDs and isolation precautions for infection control have been linked to increased rates of adverse events and higher mortality.6 34 Second, bedspacing may reduce the frequency and quality of communication between CTU medical personnel and unfamiliar allied health team members (especially nurses) who host these patients on off-service wards. Inadequate communication between nurses and physicians has been associated with higher inpatient mortality.35 36 On the GIM wards, daily ‘bullet rounds’ typically occur Monday to Friday. At these rounds, the medical and allied healthcare teams discuss and update patient care plans daily. An interdisciplinary approach has been shown to decrease mortality.37 These bullet rounds would omit bedspaced patients, and therefore these patients would not receive the benefits of a comprehensive team approach. Third, the allied health team members on off-service wards may provide lower quality of care due to different skill sets and lack of experience caring for GIM inpatients. Inexperience with managing specialised medical issues, such as care for haemodialysis patients on surgical wards, has led to care lapses that predispose patients to adverse events.10 Lastly, for an acutely sick GIM patient, the highest workload in terms of new investigations and clinical instability are often compressed into the first few days post admission. Bedspaced GIM patients are often bedspaced at the start of admission. Less frequent clinical monitoring by the team due to geographical barriers at this vulnerable period may increase the risk of adverse outcomes. We hypothesise that a combination of these factors may have contributed to the observed increase in mortality for bedspaced patients. The magnitude of the risk is greater than we had anticipated based on the presumed mechanisms, and there may be a complex interplay of various factors contributing to this striking increase in mortality. The limitations of this single centre study such as residual confounding or other unforeseen factors may have contributed to this surprisingly large mortality risk. Pending further study at other institutions, cautious interpretation of our results is warranted.

Our study showed that the highest risk of mortality for bedspaced patients occurred within the first week of admission. We hypothesise that internal medicine patients were most medically active within the first days of admission requiring intensive investigations and management. Bedspacing may have the most impact during this critical period of active workup and management resulting in a front-loaded mortality risk.

Interestingly, based on the competing risk model, the bedspaced patients had a slightly higher rate of discharge compared with assigned GIM patients initially at time of admission. However, this trend was not statistically significant for the propensity-matched and pair-matched analyses. It may be that the knowledge of taking up a bed from another service gave the CTU teams additional motivation to address the bedspaced patients’ medical issues and discharge these patients sooner.

This study highlights the importance of managing patient flow and bed mapping. Future interventional or quasi-experimental studies may help demonstrate the impact of patient flow and bed mapping policies on patient quality of care and outcomes. The number of bedspaced off-service ward GIM patients may be a clinically important and potentially modifiable quality marker. Future hospital policies focusing on patient flow should prioritise getting patients to the appropriate assigned ward beds, which may contribute to improved patient outcomes.

Our findings should be interpreted with caution, because they are based on a single centre. Each hospital has unique features relating to hospital capacity, bed mapping, admission processes, organisation of healthcare teams and training of healthcare professionals. Small single centred studies tend to overinflate the estimated effect due to a greater risk of bias.38 39 It is not uncommon for results of non-randomised studies to be contradicted by future studies.40 Therefore, future studies with multiple sites including both academic and community hospitals are necessary to ensure the generalisability of these findings. Future studies should also examine reasons for the differences in mortality, to help design strategies to minimise risk when bedspacing is deemed unavoidable.

In conclusion, we observed substantially higher inhospital mortality for GIM patients bedspaced to off-service wards. Future studies in other hospitals should help clarify mechanisms that influence the association between bedspacing and adverse outcomes including mortality. In this way, we may be able to identify modifiable contributors to this relationship and thereby improve patient safety.

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Contributors ADB and SSG designed the study. ADB acquired the data. ADB and GAT performed the analyses. All authors were involved in interpretation of data. ADB wrote the first
draft of the manuscript. All authors were involved in revisions and approving the final manuscript for publication. ADB is guarantor for the manuscript.

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**REFERENCES**


